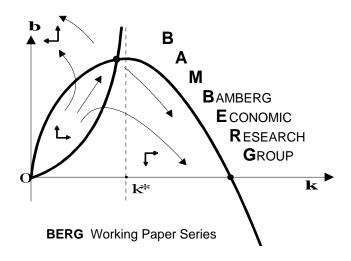
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Does the Transition into Daylight Saving Time Affect Students' Performance?

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Abstract

We use international student assessment data on more than 22,000 students from six European countries and a regression discontinuity design to investigate whether the transition into daylight saving time (DST) affects elementary students' test performance in the week after the time change. We do not find reliable statistical effects on students' performance, neither in math, science nor reading. Our results therefore challenge the prevailing public opinion that DST should be abandoned because of its detrimental effects on school children's performance.

JEL: D04, H41, I20, I29

Keywords: Daylight saving time; school achievement tests; cognitive performance; natural experiment; regression discontinuity design; TIMSS; PIRLS

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1 Introduction

80 countries around the world¹ are currently exposed to a shift in sleep patterns twice a year when they switch between daylight saving time (DST) and standard time (ST): In the northern hemisphere, clocks are set forward by one hour in spring to DST and set backward by one hour in fall to ST. While the phase delay in fall rewards us with an additional hour of sleep, the phase advance in spring implies that we have to get up one hour earlier – while the sunlight lags one hour behind.

Ever since its first introduction, the change to DST has been critically discussed. Germany and Austria-Hungary introduced DST in 1916² in order to save energy and to better match sleep-wake cycles with daylight times. Various recent studies challenge that DST saves energy (e.g. Kellogg and Wolff, 2008; Aries and Newsham, 2008; Kotchen and Grant, 2011; Sexton and Beatty, 2014). Another strand of research discusses whether the shift to DST increases traffic and work-related accidents (Hicks et al., 1983; Barnes and Wagner, 2009) or not (e.g. Ferguson et al., 1995; Lahti et al., 2011), influences stock market returns (e.g. Kamstra et al., 2000) or not (e.g. Gregory-Allen et al., 2010), affects individuals' subjective well-being negatively (Kountouris and Remoundou, 2014; Kuehnle and Wunder, 2015) or might even increase the risk of heart attacks (e.g. Janszky et al., 2012) or not (e.g. Sandhu et al., 2014), to name just some. So far, no clear conclusions can be drawn as to whether DST is indeed harmful enough to affect outcomes measurably and whether its potential costs outweigh its supposed benefits. Most previous studies suffer from small sample sizes (as already noted by Gregory-Allen et al., 2010) or fail to control for unobserved structural differences before and after the time change or between DSTand non-DST-countries. Nevertheless, a recent representative survey puts the share of DST-opponents in the German population at nearly three quarters (forsa Gesellschaft für Sozialforschung und statistische Analysen mbH, 2015) and there, as well as in many

¹ Data compiled from the CIA World Factbook Central Intelligence Agency (2013).

Reichsgesetzblatt (RGBI) (1916). Bekanntmachung über die Vorverlegung der Stunden während der Zeit vom 1. Mai bis 30. September 1916, RGBl 1916, p. 243.

other countries, regular petitions urge parliaments to break with the tradition of changing clocks twice a year.

In addition, and although one can set one's clocks to reading regularly in the newspapers that DST should be abandoned because it is detrimental to school children's performance, there is hardly any scientific evidence on whether the time change affects school performance. To the best of our knowledge, there is only one study (Gaski and Sagarin, 2011) on the long-run impact of the semiannual clock changes and students' performance in the Scholastic Aptitude Test (SAT) in Indiana, USA. The authors report 16% of a standard deviation lower SAT scores in DST-adopting counties. Yet, we doubt that the difference the authors find can be explained by the clock change: Most U.S. states experience changes of more than 10 points in mean SAT scores in reading and math over time (National Center for Education Statistics, 2013), half of the SAT-taking schools experience a rise or fall in scores by 10 points every year, and about 20% tend to have 20 points higher or lower test scores when compared to the previous year (College Board, 2014). Moreover, Gaski and Sagarin (2011) do not account for potential structural differences between counties that might drive the results.

This paper is the first to study whether the clock advance induces short-run consequences on students' performance in six European states. In a regression discontinuity design, we exploit the fact that several countries collected data for the international student assessments *Trends in International Mathematics and Science Study* (TIMSS) and *Progress in International Reading Literacy Study* (PIRLS) during the transition from ST to DST in spring 2011. This approach provides us with a sample of more than 22,000 students. Hypothesizing that elementary school children might suffer from sleep deprivation in the week after the switch to DST, we investigate whether moving the clock forward by one hour affects students' performance.

This mechanism is backed up by a rich literature on the relationship between sleep and performance, which indicates that cumulative or complete sleep deprivation decreases cognitive test performance (e.g. Astill et al., 2012; Van Dongen et al., 2003; Banks and Dinges, 2007; Goel et al., 2009). However, it is not known whether the rather mild and short-term disturbances of the circadian system introduced by the clock change are large enough to affect children's school performance significantly. If so, this would not only make a case for another debate on whether to abandon DST or not, but would also cast doubt on the validity of exams and international student achievement tests timed around the clock change. If the shift into DST does not cause large enough drops in students' performance, this would challenge the public assertion that children suffer from the clock change.

Our results challenge the predominant expectation that the clock change introduces strong and measurable changes in children's school performance. Although we do find small decreases in performance after the clock change in most countries for math and science, these effects are very small in magnitude and not significantly different from zero. Moreover, the treatment effects for reading are pointing to the opposite direction and are of similar magnitude, though also not statistically significant. Checking the robustness of our results with TIMSS data on eighth-graders reveals that our findings cannot be explained by the young age of the fourth-graders in our main sample.

2 Effects of the clock advance

2.1 The circadian clock

In each of us ticks a circadian clock that determines when we sleep and when we wake. Daylight serves as a zeitgeber to our inner clocks and synchronizes our sleep-wake patterns approximately (circa) to the daily (dian) rotation of the Earth. Our organism is tied to that inasmuch as the hormone melatonin regulates our sleep-wake-cycle by sending us to sleep. When it gets dark, our bodies produce melatonin and we begin to feel sleepy. At dawn, the production is stunted and we awake.

In Europe, this bio-chemical system is disturbed each spring when the clock is set forward by one hour in the very early morning hours of the last Sunday in March, and our alarm clock conflicts with our inner clock: While school or work still start at, say, 8 a.m. clock time, our bodies continue to follow the light-dark cycle that still lags one hour behind. In the mornings, melatonin levels are still up and we feel sleepy. In the evenings, we have difficulty falling asleep. This deprivation of sleep persists until the DST and the light-dark cycle are synchronized, or in other words, until we have settled the dispute between the alarm and inner clock (Valdez et al., 2003, p. 146).

2.2 Sleep, light and cognitive performance

Both sleep and light are also correlated with cognitive performance. Light does not only affect vision but exerts a direct positive effect on the functioning of the brain and its availability increases cognitive performance (Heschong et al., 2002; Vandewalle et al., 2006, 2009).

The positive association between sleep duration and cognitive test performance of adults is well documented (e.g. Van Dongen et al., 2003; Banks and Dinges, 2007; Goel et al., 2009).³ A recent meta-analysis shows that also for 5-12 years aged children, sufficient sleep is significantly related to higher cognitive performance, less internalizing (e.g. anxiety, sadness) and externalizing (e.g. aggression, hyperactive behavior) behavioral problems, and, especially, better performance in school (Astill et al., 2012, and references therein). At the same time, children's attention, memory and intelligence seem to be unaffected by sleep duration (Astill et al., 2012).

Correlational studies draw the picture of a negative relationship between self-reported hours of sleep and grades in middle and high school (consult Wolfson and Carskadon (2003) or Shochat et al. (2014) for a review). Children seem to be sensitive to small or modest changes in sleep duration. In that vein, Vriend et al. (2013) show that reducing habitual sleep duration of 32 children by one hour for four consecutive nights affected

The performance enhancing effects of sleep even seem to pay off in monetary terms. Instrumenting sleep duration with the local sunset time, Gibson and Shrader (2014) estimate the causal effect of hours slept on wages. Speculating that an earlier sunset drives people to bed earlier, the authors provide evidence that sleeping one hour more each night increases wages by 16%.

children's mood and emotional regulation negatively and decreased their cognitive performance.

2.3 Sleep and performance after the clock change

It is even less clear, in how far all these processes carry over to the clock change, especially with respect to elementary school children who are the subject of this study.

First, there is mixed evidence on how long the sleep-wake cycle needs to adapt after the clock change. Results from both early and recent studies indicate that children and adolescents lose between 40 and 50 minutes of sleep following the switch from ST to DST (Reese, 1932; Barnes and Wagner, 2009). Schneider and Randler (2009) report that school children showed a higher daytime sleepiness after the time change. The adaption process to the new regime can take up to several weeks, depending on chronotype and sleep patterns during weekends (Valdez et al., 2003; Schneider and Randler, 2009). In contrast, adjustments to phase delays as encountered when clocks are reset to ST, traveling westwards or moving from daytime shift work to night shift work are easier and faster (Hauty and Adams, 1965a,b; Lemmer et al., 2002; Niu et al., 2011).

Second, the impact of a single small short-term shift in the circadian clock is only rarely studied. For instance, Burgess et al. (2013) simulated small disturbances of the circadian clock in 11 adults, who reacted with significantly slower reaction times in a Psychomotor Vigilance Test. Monk and Aplin (1980) analyzed the performance of 39 adults during the shift from DST to ST, i.e. during the phase delay in fall. After waking under the standard clock time, subjects showed enhanced performance in calculation tests. Yet, the authors could not separate this effect from the simultaneous effect of a better mood on awakening.

Recently, three economic studies provided quasi-experimental evidence of a later school start time on student achievements. Although these studies can rely on larger samples and exogenous variation instead of self-reported measures from survey data, the results are again mixed. Edwards (2012) exploits the fact that U.S. middle schools

start the school day at different times to reduce the costs of the public transportation system. Using between and within variation, he finds a 2-3 percentage point increase in standardized math and reading test scores when school starts one hour later. Carrell et al. (2011) show that delaying course start times by 50 minutes increased students' achievements at a US-military post-secondary institution by as much as a one standard deviation increase in teacher quality. The authors use variation from two sources: First, starting times were shifted step-wise from 7:00 to 7:30 and finally to 7:50 AM. Secondly, some students were randomly allocated to early courses only, others to later courses only and the rest attended both early and late courses. Note however that students allocated to later courses could not use the additional time in the mornings to sleep longer but were required to attend the early breakfast with their fellow students. Carrell et al. (2011) argue that late-starting students could have taken a nap between breakfast and their first class, thereby getting more sleep and performing better throughout the day. Given that the military institution prohibited napping (p. 78), it is contradictory that additional sleep should be the main driver of higher performance. To us, it seems equally likely that students in late courses achieved higher grades because the empty time-slot allowed them to repeat and thereby better remember the course content. This could also explain why the treatment estimates of attending an early class loose their statistical significance once student fixed effects are included. In contrast to that, Hinrichs (2011) uses longitudinal individual data on the U.S. high school achievement test ACT and exploits, in his main analysis, exogenous variation from a policy change in the U.S.: While Minneapolis and some of its surrounding districts shifted school starting hours backwards, its Twin City St. Paul and surroundings retained the old starting times. The author does not find evidence for the hypothesis that ringing the school bell later increased students' performance.

To the best of our knowledge, there is only one study on the relationship between DST and performance of students. Using the variation in DST-regimes between counties of the U.S. State of Indiana, Gaski and Sagarin (2011) identify the long-run effects of DST

on county-wide SAT test performance. The authors find test results to be significantly worse in counties that advance and set back their clocks each year when compared to counties sticking to ST permanently.

Note that our approach, outlined in the following section, is different. Gaski and Sagarin (2011) compare long-run average performance in counties that do or do not change their clocks. Their approach comes at the risk of mistakenly interpreting structural differences between counties as causal effects. The authors do, for instance, not control for the proximity to large cities outside Indiana. It seems plausible that the counties close to Chicago, Cincinnati or Louisville change the clocks to synchronize working times for commuters from Indiana. Worse SAT scores could then e.g. be due to the reduced time commuting parents and their children spend at home together or a less privileged background also explaining why they cannot afford living closer to the city. In contrast to that, our study focuses on short-run effects of the clock change within DST-adapting countries. Exploiting the random allocation of schools to test dates before and after the clock change as a natural experiment allows us to separate the effect of the transition into DST from structural or institutional differences. If the mild disturbance of the inner clock affects sleep patterns so much that performance in the week after the change suffers, we should be able to observe a short-run dip in performance.

3 Method

We analyze the shift to DST as a natural experiment to study before-after differences in students' performance. As sleep-wake cycles and human performance are thought to synchronize within about one week after the clock change (Valdez et al., 2003), we restrict the sample to schools tested within one week before and one week after the change to DST.

More specifically, we regress the test score TS_{ijc} of student i in school j and country c on the treatment indicator, DST_{ijc} , a set of controls, x_{ijc} and a the constant η_0 . α_0 is

the coefficient of interest as it captures the effect of the switch to DST on student test scores. We run hierarchical linear models (HLM) with maximum likelihood to account for the nested structure of the data.⁴ The error term is therefore a composite taking care of the different variance between schools, v_j , countries, v_c , and the remaining individual errors, ϵ_{ijc} :

$$TS_{ijc} = \eta_0 + \alpha_0 \cdot DST_{ijc} + \boldsymbol{x}'_{ijc} \cdot \boldsymbol{\beta} + v_j + \nu_c + \epsilon_{ijc}. \tag{1}$$

As students are only tested once - either before or after the clock change - our regression discontinuity design relies on the identification assumption that the assignment to test dates before (control group) and after (treatment group) the shift to DST was random. Restricted to a given time span determined by the end-dates of the school year, the TIMSS and PIRLS testing dates are agreed upon by the school coordinators and the testing agencies. Given sufficient capacity on the test agency's side, the students' performance was assessed on that day.⁵ The sampling is therefore unrelated to regional characteristics (south/west, rural/urban) that might have also driven the test score results. 6 This procedure might however open up the possibility of self-selection into treatment and control group. For example, if coordinators of good schools preferred test dates before the shift to DST because they anticipate a dip in their students' performance, we may mistakenly contribute a negative treatment effect to the clock change, while it only captures a generally worse performance of students tested later. Although we cannot fully rule out that consideration of the clock change mattered when school coordinators proposed a testing date, we consider it unlikely that coordinators were aware of the clock change and its potential harmful effect on their students' performance as dates were scheduled well in

We also estimated OLS-models with standard errors clustered on the highest level, i.e. on the school level in the country-specific regressions or the country level in the pooled sample. The coefficients estimated with OLS were similar.

Most schools were tested only on one day per study. In 3.57% of TIMSS- and 3.35% of PIRLS-schools, a few students were tested after the clock change although their school was sampled before the clock change – probably because they were ill during the main testing time and data for the missing students was collected later.

⁶ A systematic geographical sampling would have introduced the risk of mistaking structural differences or differences in the availability of daylight between eastern and western areas within a country for a performance difference with respect to the DST-shift.

advance.⁷ Moreover, assessments took place at the end of the respective school terms, i.e. during a period where schools schedule examination board meetings, field days or other activities filling the students' and teachers' timetables. We do therefore expect that it is challenging enough to arrange a test date that fits the students', teachers' and testing agencies' schedule without consideration of the clock change. Apart from that, it is impossible to identify single schools in the data later, reducing any possible incentive for school coordinators to optimize their students' performance with respect to test time selection – given that they were indeed aware of the clock change date.

We check the plausibility of the identifying assumption by comparing treatment and control group students on variables that might drive test performance. To do this, we test whether potential differences in means of all covariates between treatment and control group are significantly different from zero. To account for the fact that very small differences between treatment and control group lead to high values for the t-statistic if the sample size is large, we also calculate the scale-free normalized differences as suggested by Imbens and Wooldridge (2009, p. 24). More specifically, we take the differences in means between covariates before, x_{before} , and after, x_{after} , the treatment and normalize them by their sample standard deviations, using the respective sample variances before, s_{before}^2 , and after, s_{after}^2 , the treatment:

$$\Delta_x = \frac{x_{after} - x_{before}}{\sqrt{s_{before}^2 + s_{after}^2}}.$$
 (2)

Following the authors' rule of thumb, we interpret differences larger than a quarter of a standard deviation as indication of selection bias and sensitivity of linear regression with respect to model specification.

To account for potential differences in performance over the week, e.g. a "blue Monday effect" or exhaustion over the week (Laird, 1925; Guérin et al., 1993), but also to in-

According to the National Research Coordinators of the five TIMSS countries in our sample, the majority of schools was first contacted some 6 to 8 months prior to the testing day.

vestigate whether the DST-effect fades out over the week, we include control variables for each testing day of the week, day, and its interaction with the treatment indicator.

$$TS_{ijc} = \eta_1 + \alpha_1 \cdot DST_{ijc} + \sum_{d=2}^{5} \gamma_d \cdot day_{id} + \sum_{d=2}^{5} \delta_d \cdot day_{id} \cdot DST_{ijc} + \boldsymbol{x}'_{ijc} \cdot \boldsymbol{\beta} + \nu_j + \nu_c + \epsilon_{ijc}.$$
(3)

We use Monday as the reference category. Therefore, α_1 represents the treatment effect for Mondays after the treatment. The marginal effect of the time change on the Tuesday under DST equals then, for instance, the treatment dummy, α_1 , plus the coefficient of the treatment interacted with the performance on d=2 after the clock change, δ_2 .

4 Data

The International Association for the Evaluation of Educational Achievement (IEA) assesses fourth- and eighth-graders' reproduction, application and problem solving skills in several areas of math and science since 1995 in the the Trends in International Mathematics and Science Study. The Progress in International Reading Literacy Study is conducted every five years and measures trends in fourth-graders' reading literacy and comprehension.

We can make use of several fortunate coincidences in the latest currently available wave of 2011 which we use for the following analyses: First, 2011 is the only year for which we can use assessment data for all three testing areas (math, science and reading) because both TIMSS and PIRLS data were collected. Secondly, while data on the exact date of the testing was not contained in previous waves, this information is available in the 2011 waves. Lastly, as student achievement data were collected in the last weeks of the respective countries' school terms, the field phases of several countries coincided with the transition into DST. In TIMSS 2011, there was an overlap between fourth-graders' testing dates and the clock change in seven countries. Being especially interested in performance differences on the Monday after the clock change (which was March 28, 2011), we had to exclude the two countries that lacked test data on Mon-

days (Finland and Ireland). After excluding four students who were tested on a Sunday and 309 cases with missings on our covariates, our analytic sample from TIMSS contains 8,813 fourth-graders in 364 schools from Denmark, Lithuania, Norway, Spain and Sweden.⁸ As Denmark did not participate in PIRLS 2011, our PIRLS sample includes Lithuania, Norway, Sweden and Spain, but also Finland where students' reading performance was assessed on all weekdays. After listwise deletion of 357 cases, our analytic PIRLS sample sums up to 13,255 fourth-grade students clustered in 508 schools.

TIMSS and PIRLS follow a matrix-sampling approach, meaning that there are many more questions than answered by a single student in their assessment booklets. Whereas students answer only one booklet, each item is questioned in more than one booklet. This overlap is used to construct an estimate of the achievements in the student population with the help of scaling methods from item-response theory (Mullis et al., 2009b,a; Yamamoto and Kulick, 2012, p. 123). To account for the uncertainty introduced by imputing the scores, the International Association for the Evaluation of Educational Achievement provides five *plausible values* for the achievement of each student. We retain this uncertainty by using all five plausible values in the following analyses. Achievement scales range usually from 300 to 700 points. To establish comparability over time and between countries, achievement test scores were scaled in 1995 (TIMSS) and 2001 (PIRLS) to an international mean of 500 and a standard deviation of 100. International benchmarks to classify students' achievements into low, intermediate, high and advanced levels were set at 400, 475, 550 and 625 points.

Note that for the following short description of our data, we focus on TIMSS in order to save space. We provide statistics for our PIRLS sample in the appendix and outline only main points in this section. Table 1 gives an overview of the plausible values for the countries in our sample, showing that students perform between the intermediate and

We focus on students in grade 4 as the data for eighth-graders do only include two countries (Sweden and Finland) for the respective time period and reading literacy is not assessed in grade 8. We do, however, draw on the eighth-graders sample in our robustness checks.

More specifically, we apply appropriate estimation tools as implemented in the multiple imputation (mi) commands in Stata for all analyses.

the high international benchmark. On average, students achieve a score of about 509 points in math (S.D. \approx 72) and 515 points in science (S.D. \approx 70). Spanish students score lowest and Danish students highest in math. In the science assessment, test scores are highest for Swedish and lowest for Norwegian children. For PIRLS, we find an average of 537 points in reading (S.D. \approx 68), and that the best readers in our sample are the Finnish students, while the elementary school children in Norway achieve the lowest scores (cf. appendix, table A1). Comparing test scores before and after the time change on this descriptive basis already suggests that there are no large differences before and after the transition into DST.

- Table 1 about here -

Table 2 contains further descriptive statistics on our later controls. We include gender and age in months to investigate heterogeneous effects and to account for potential differences between treated and controls. All of our students are in grade 4 and the average student is about 10 years old. Half of the sample is female. The high performing Danish, Finnish and Swedish students are, on average, one year older than the lower performing Norwegian and Spanish children.

- Table 2 about here -

We include an indicator for whether students wrote the test in the language they speak at home to control for language-related differences in test scores. About 80% do indeed always stick to the test language at home and only 2-3% indicate to never use it at home. To control for the children's socio-economic background by proxy, we add the number of books at home. Most children indicate that their parents have 26-100 books (one bookcase) at home. In 14% of the cases, children report more than two bookcases

There are several reasons why we decided in favor of this often used proxy for the educational, social and economic background of the family. Firstly, the number of books at home is easily comparable across countries (Wößmann, 2004). Secondly, the predictive power of the books variable with respect to student performance is higher than that of parents' educational background (Wößmann, 2003). Lastly, while books at home are reported for nearly all students, parents' educational achievement is systematically missing for about one third of the cases in our TIMSS sample and about 12% in our PIRLS sample. Missing cases are a selective sample of students with a low number of books at home.

(more than 200 books) at home. The average number of books at home is relatively high in Sweden and Norway and relatively low in Lithuania.

When turning to test days, the table shows that most students were tested on Tuesdays or Wednesdays. In our TIMSS-sample, 14% of the overall sample was tested on Mondays (table 2), thereof 43% before and 57% after the switch to DST. 15% of the PIRLS-students were tested on a Monday (table A2), 28% of them under ST and 72% under DST.

As outlined in the previous section, we test for (normalized) differences between students treated before and after the clock change. The results are reported in table 3 for TIMSS and table A3 for PIRLS. While absolute differences are statistically significant for most variables, they show neither a systematic pattern nor are the normalized differences above the critical value of 0.25 suggested by Imbens and Wooldridge (2009, p. 24). Including these covariates in the following regressions controls for slight differences between groups that should not substantially affect our results.¹¹

- Table 3 about here -

5 Results

5.1 Performance-effects of the clock change in the pooled sample

Table 4 reports the effects of the clock change on students' performance in math, science and reading for the pooled sample of all countries. Note that the students who where tested in math were also tested in science and vice versa, because both fields were part of the TIMSS study. Most of the TIMSS students did also participate in the PIRLS reading assessment, but not all of them.¹²

We also investigated differences in other proxies for the students' socio-economic status, e.g. own possessions including books, study desks or computers. We did not find a systematic pattern within and over countries that would speak for a selection of specific students/schools to the treatment or control group. Moreover, our results were very similar after including these variables as additional controls.

We would have liked to also present within-analyses for students who participated in both TIMSS and PIRLS and completed one study before and the other one after the clock change. Unfortunately, TIMSS and PIRLS were always conducted at consecutive days within the same week.

- Table 4 about here -

In the HLM specification without covariates (table 4, column 1), students scored about 4 points lower in both math and science when tested during the week after the clock change. Given a standard deviation of about 72 and 70, test scores in the week after the clock change drop by circa 6% of a standard deviation. Yet, these effects are neither substantial in terms of statistical significance nor in terms of magnitude. This gets even clearer when looking at the estimate for reading, indicating that students performed less than 1% of a standard deviation better when sampled after the clock change. Again, the effect is not statistically different from zero.

Two points stand out after including weekday dummies and weekday-treatment interaction terms as of equation 3 in column 2. Firstly, students' pattern of performance in the week before the time shift is rather stable (figure 1). Secondly, as the Monday before the time change is our reference category, the treatment coefficient in column 2 shows the reduction in students' test scores at the Monday immediately after the switch to DST. We would expect the treatment effects to be largest for the first day of the week when only one night passed since the clock had been advanced. For both math and science, the coefficients imply that the negative effect on students' test scores would indeed be strongest on Mondays after the shift where students might suffer most from sleep deprivation. Contradictory to that, PIRLS students showed a slightly better reading performance on the Monday after the clock change than on average over the week. However, as the standard errors increase by about the same rate, all point estimates remain statistically insignificant. Plotting the average performance levels shows that students perform only slightly worse after the clock change as can be seen from the dashed line in figure 1 for math (panel a) and science (panel b). For reading, we observe a very small

To the best of our knowledge, there are only two studies on day-of-week effects in students' performance with contradictory results. Laird (1925) finds students performance highest on Wednesdays. While Laird (1925) uses a multi-faceted measure of cognitive abilities, Guérin et al. (1993) test 8- to 10-year old girls' attention, mental speed and visual scanning abilities in a letter cancellation test. They do not find a pattern for 8-year-olds but do find peaks in performance for the 10-year-old girls on Tuesdays or Fridays.

positive effect (panel c). Please keep in mind, however, that these differences are always small and not statistically significant.

- Figure 1 about here -

Adding covariates for gender, age, age squared, books at home and whether test language is spoken at home instead (column 3) does only slightly decrease the estimates of the treatment effects when compared to column 1, confirming our identification assumption. As given in full detail in the appendix, tables A4 to A6, the signs of all covariates moreover follow expected patterns with female students scoring lower in math and science but higher in reading, and diminishing positive effects of students' age on test scores. ¹⁴ In the TIMSS-sample, students who never talk in the test language when at home score at least one third of a standard deviation below those who always speak it. Obviously, this effect is even larger for reading abilities where children who never used the test language at home score half a standard deviation lower than children who took the test in their mother tongue. The coefficients for books at home also show the expected signs: Students with low socio-economic status and up to 10 books at home score nearly 75% of a standard deviation lower in science than those with a socio-economic status close to the sample average. The respective effects in math and reading are a bit smaller (68% and 63% of a standard deviation).

Considering the full model (column 4) where *DST* needs to be interpreted as the treatment effect on Mondays after the shift yields slightly different estimates for reading but not for math and science when compared to column 2.

We test whether girls' or boys' performance is more sensitive to sleep deprivation in table 5, columns 1-2. We find no gender effects for TIMSS-students, which is in line with Monk and Aplin (1980).

Children's sleep deprivation might be dependent on their socio-economic background if the latter is correlated with factors that do also determine the children's sleep duration,

Older students achieve better test score results as we already noted on the descriptive level. As higher age is, however, also a sign of grade repetition or late school enrollment due to possible developmental delay of the child, test scores do not proportionally rise with age.

e.g. parents' weekend work. We report heterogeneous treatment effects by students' confidence with the test language (table 5) and the number of books at home (table 6).¹⁵ Although students who report a very large number of books on their parents' shelves were most strongly affected on Mondays after the clock change in the TIMSS-sample, this effect is still far from being significantly different from zero and not backed up by the PIRLS-sample. We also do not find any indication of heterogeneous effects when we split the sample by whether the test language was spoken at home.

- Tables 5 and 6 about here -

5.2 Performance-effects of the clock change in the country-specific samples

Table 7 presents the treatment effect estimates by countries. It can be seen that the pattern we described for the pooled sample is also reflected in most of the country-specific samples. The effects on students' performance is strongest in Norway and Sweden. Norwegian and Swedish students score about one third of a standard deviation lower in math on Mondays after the clock change. The equivalent treatment effects for science are approximately 22% of a standard deviation in Norway and a 30% of a standard deviation drop in performance in Sweden. But again, none of these effects is significantly different from zero.

- Table 7 about here -

A word of caution is in order when investigating the test results for Denmark. In Denmark, only 17 control and 64 treatment group students were tested on a Monday. What is more is that no students were tested on the Friday before the phase delay. Therefore, we do not only lack power to identify any significant effect for Mondays and Fridays but get

¹⁵ For these and all other following estimations, we only give the treatment effects in order to save space. Full results are available upon request.

highly imprecise estimates for the treatment effect which might explain the rather large positive though, again, insignificant rise in test scores after the transition into DST.¹⁶

The results for reading are equally mixed and estimates show even positive signs in Lithuania, Sweden and Spain, although all treatment effects are not significantly different from zero.

5.3 Extensions and robustness checks

5.3.1 Two weeks before and after the clock change

If the insignificant decreases in performance in the week after the clock change were just noisy deviations from zero, we would expect that our estimates of the test results two weeks after the clock change are also not clear-cut. If, to the contrary, our one-week effects are factual decreases in performance that are not harmful enough to become significant, we would expect one of the following patterns for treatment effects two weeks before and after the time shift: Either the mildly disrupted circadian clocks have synchronized to the light-dark cycle and two-week treatment effects are consistently and remarkably smaller than the one-week effects. Or students' sleep deficits have accumulated because they did not adapt their bed hours to the new system. In the latter case, we would expect systematically larger and potentially significantly negative treatment effects.

Considering the longer observation period of two weeks before and after the clock change moves most of our estimates of the average performance in the two weeks after the clock change closer to zero, but only very slightly and not consistently (table 8). When we break down the analysis to the five countries in our TIMSS-sample and compare the already discussed performance estimates on Mondays immediately after the clock change (table 7) with our new estimates for the 2-week-window (table 9), we

Moreover, Lithuania sampled only 93% of the international target population, namely those students taught in Lithuanian (Joncas, 2012), whereas all other countries in our sample did not impose a restriction with respect to language of instruction. This sample selection might be one reason why the Lithuanian treatment effects are slightly different.

do also not find a clear trend: The average TIMSS-performance of the two Mondays after the time change is lower in Denmark and Sweden (in the latter does the estimate even turn significantly negative), but higher in Lithuania. This ambiguity is also mirrored in the PIRLS-sample.

- Tables 8 and 9 about here -

5.3.2 Age effects: eighth-graders

The fact that we do not find negative effects on performance in school might be due to the young age of the children in our sample. As children need more sleep in general and are rather morning-chronotypes getting sleepy early in the evenings, they might have less trouble falling asleep when sent to bed earlier (Valdez et al., 2014) and recover fast from a sleep deficit. If parents anticipate the clock change and slowly familiarize their children to the new time regime in the days before the switch to DST, the sleep deprivation of young children after the time change would be minimal. Adolescents are not as easily convinced to go to bed one hour earlier. Moreover, during puberty, chronotypes shift to evening types and adolescents have trouble to go to bed early and to rise early in the mornings (Valdez et al., 2014). Edwards (2012) indeed shows that positive performance effects of delaying school starting time increase with age. For elementary grade students, the author does not find an effect, probably because young children are not exposed to changes in their hormone regulation which also affect sleep-wake-patterns. However, he cannot rule out that his results are driven by later school starting times in elementary school.

To shed some light on whether the impact of the clock change depends on age and whether our insignificant results are driven by the fact that we use a sample of 10-year-olds, we repeated all analyses with 2011-TIMSS-data on eighth-graders who are right in the middle of puberty (about 15 years old). Only Finland and Sweden sampled 5.591 students in 199 schools within the time period we are interested in. Descriptive statistics

show that the Finnish and Swedish students in the sample are very similar and mirror the descriptives of Finnish and Swedish fourth-graders we presented above.¹⁷

When we compare the size of the treatment effects for eighth-graders in math and science as reported in table 10 with the effects of fourth-graders in the respective fields, we see that the effects for Finnish eighth-graders are not larger: Finnish students score insignificant 2% of a standard deviation lower in math and 3.65% of a standard deviation higher in science on the Monday after the clock change. Two weeks after the time-shift, both effects are slightly negative but again not statistically significantly different from zero. In Sweden, we find negative effects of about one third of a standard deviation (table 10). The magnitude of this treatment effect is rather large and comparable to the difference between students of families with one bookshelf and those with one bookcase at home. The effect is also marginally statistically significant in math and significant at the 5%-level for science. These effects drop slightly in the 2-week-specification but only the science coefficient stays significant on the 10%-level. Note however that, although we find the effects for Swedish 15-year-olds now statistically significantly different from zero, the estimates are not larger than the dips in the fourth-graders' performance. Our results do therefore not confirm the hypothesis that DST is more harmful to older students.

- Table 10 about here -

5.3.3 Testing times

Unfortunately, we lack data on the exact times of the assessment, except for Lithuania where the National Research Coordinator provided us with additional information. While the official IEA instruction for schools required to take the tests in the morning hours, we cannot rule out that schools conducted the tests at different times or had different school starting times. As cognitive performance typically increases over the morning hours (Vandewalle et al., 2009), students who took the test later might have

Again, we do not present tables for most of our age effects analyses. All tables are available upon request.

performed better. However, this would only affect our results if the schools sampled after the clock change had started the assessment systematically later. Yet, at least in the Lithuanian case, we did not find an indication of assessments being scheduled to later time slots after the clock was advanced by one hour.

6 Discussion

In this paper, we investigated whether the transition into daylight saving time in spring and a potentially associated sleep loss in the following nights affect students' performance in international assessment tests. Our findings challenge the prevalent public feeling that the clock change harms school children's performance. We exploited the fact that schools participating in the large-scale international student assessment studies TIMSS and PIRLS were randomly allocated to testing times before and after the time shift. In a regression discontinuity set-up, we did not find elementary school children to be measurably affected by the clock change. The estimated treatment effects were most consistent with the hypothesis of a sleep deprivation effect in Norway and Sweden. But even for these countries we failed to find statistically significant effects. For reading and the other countries in our sample, we did even find positive, yet insignificant, effects.

To investigate several further hypotheses and possible explanations for our findings, we varied the time window around the clock change. Having found no indication of an accumulating or decreasing sleep deficit in elementary school children, we conclude that these young children did not suffer from a harmful sleep deficit. On the one hand, this might be due to the fact that one hour of sleep loss is not enough to disbalance circadian clocks by so much that performance within the following week suffers measurably. On the other hand, a lot of websites on the internet provide hints on how to adjust childrens' circadian systems smoothly to the new schedule, e.g. by preponing bedtimes by some minutes each day before the transition into DST. However, we wonder whether parents do and can indeed plan their fourth-graders sleep pattern so diligently.

If the latter explanation were however true or if the harmfulness of DST were in general depending on the students' age, we would expect higher treatment effects in a sample of older students. We briefly explored this hypothesis, using a TIMSS sample of 15-year-olds, and did not find larger treatment effects. However, the treatment effect for Swedish children was not only rather large but also statistically significantly different from zero. This effect might point the Swedish children's higher sensitivity to yet another phase change introduced by DST. In the north, the periods of darkness are very long in winter and very short in summer. Swedish students' circadian rhythms are therefore more often affected by changes in the light-dark cycle. Previous studies found long nights to increase the prevalence of seasonal affective disorders (Rosen et al., 1990) and long days to negatively affect mental health and even increasing suicide rates (Björkstén et al., 2009). It is however questionable why we did not find large and significant effects for Finland in the eighth-graders sample as well. Part of the reason might be the fact that Finland is further to the east and students might profit from more sunlight in the mornings when they are on their way to school. Although we know that school must have started after dawn for all students in all countries in the first week under DST, the duration of exposure to sunlight might be important as it regulates melatonin – and therefore sleepiness – levels. But even in the Swedish case, we are cautious to interpret our results as indicative of DST introducing permanent disruptions of the circadian system as Gaski and Sagarin (2011) do for the U.S. because many previous studies show recovery effects in sleep patterns and adaption of the circadian clocks to DST. We do rather believe that the shift into DST is no problem for South European or mid-European students but might be an additional challenge for students in the northwest where long nights and days are already disturbing the rhythm of our inner clocks. As we lack geographic identifiers for single schools in our data, we cannot further investigate these hypotheses but consider it as a fruitful analysis for prospective studies. It will also be an issue for future work to test the robustness of our age effects in a richer data set of more countries and students at many different ages.

Based on our research, it is however fair to say that neither parents nor children nor competence testing agencies (or even policy makers) have reason to worry about allegedly harmful effects of the transition into daylight saving time.

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Tables and figures

Table 1: Plausible values by country and treatment

| | Po(| Pooled | Deni | Denmark | Lith | Lithuania | Nor | Norway | Spain | ain | Swe | Sweden |
|---------------|---------|---------|---------|---------|---------|-----------|---------|---------|---------------|---------|---------|---------|
| | before | after | before | after | before | after | before | after | before | after | before | after |
| Math score | | | | | | | | | | | | |
| – 1st PV | 511.86 | 507.07 | 542.09 | 532.31 | 534.23 | 538.27 | 498.80 | 495.60 | 495.60 496.29 | 488.65 | 508.85 | 499.36 |
| | (71.19) | (72.43) | (67.89) | (73.12) | (74.33) | (70.09) | (68.84) | (68.40) | (70.06) | (68.85) | (66.06) | (68.80) |
| - 2nd PV | 511.34 | -, | 542.24 | 530.41 | 532.89 | | | | 495.38 | | 509.40 | 497.27 |
| | (71.70) | (72.70) | (66.72) | (74.83) | (74.13) | (70.67) | (69.65) | (68.76) | (70.91) | (68.59) | (67.17) | (67.38) |
| – 3rd PV | 511.82 | 506.72 | 540.56 | | | 539.65 | 497.36 | | 496.22 | 487.59 | | 497.07 |
| | (70.43) | (72.46) | (60.53) | (72.71) | (72.84) | (69.95) | (67.94) | (68.40) | (71.06) | (68.55) | (66.42) | (67.54) |
| – 4th PV | 512.08 | 507.17 | 541.87 | 533.10 | 535.65 | 540.17 | 497.76 | 495.07 | 495.60 | 488.47 | 510.11 | 496.79 |
| | (71.16) | (72.86) | (65.44) | (70.97) | (73.01) | (70.18) | (68.71) | (69.91) | (70.04) | (68.19) | (67.24) | (69.50) |
| – 5th PV | 511.44 | 507.35 | 540.29 | 533.40 | 535.38 | 538.92 | 497.38 | 496.31 | 495.08 | 488.61 | 509.15 | 498.30 |
| | (71.19) | (72.48) | (64.52) | (71.16) | (72.93) | (70.67) | (68.80) | (68.69) | (70.56) | (68.53) | (67.36) | (68.37) |
| Science score | | | | | | | | | | | | |
| – 1st PV | 519.91 | 513.55 | 535.04 | 519.19 | 516.16 | 522.16 | 498.82 | 495.09 | 522.26 | 515.08 | 539.08 | 527.25 |
| | (69.25) | (68.98) | (68.98) | (73.25) | (64.72) | (64.65) | (63.16) | (63.07) | (67.95) | (70.80) | (72.79) | (76.38) |
| - 2nd PV | 517.54 | 512.72 | 531.42 | 523.28 | 513.31 | 520.23 | 495.73 | 494.54 | 519.12 | 514.08 | 538.58 | 526.66 |
| | (70.79) | (09.69) | (71.30) | (75.42) | (66.33) | (65.07) | (63.25) | (64.25) | (71.69) | (70.96) | (73.70) | (77.19) |
| – 3rd PV | 517.98 | 512.26 | | | 513.12 | 521.60 | 496.70 | 495.40 | 521.20 | 512.14 | 538.08 | 523.83 |
| | (70.78) | (68.84) | (68.89) | (74.39) | (65.23) | (65.42) | (64.64) | (63.56) | (73.22) | (69.75) | (73.36) | (75.92) |
| – 4th PV | 517.89 | 512.21 | 534.20 | 524.51 | 513.97 | 521.77 | 496.73 | 494.73 | 520.19 | 511.81 | 536.96 | 523.93 |
| | (71.27) | (68.84) | (67.89) | (74.80) | (66.37) | (63.56) | (64.49) | (65.27) | (74.05) | (69.75) | (74.48) | (75.74) |
| – 5th PV | 519.69 | 513.56 | 533.04 | 525.01 | 514.72 | 519.04 | 497.95 | 496.74 | 523.21 | 514.89 | 540.28 | 529.29 |
| | (71.00) | (89.68) | (65.64) | (69.79) | (66.90) | (65.38) | (64.16) | (65.46) | (72.66) | (70.99) | (74.30) | (78.29) |
| Observations | 4000 | 4813 | 309 | 255 | 823 | 1293 | 1154 | 1174 | 601 | 1607 | 1113 | 484 |
| | | | | | | | | | | | | |

Notes: Plausible values (PV) before and after the clock change as contained in TIMSS 2011. Standard deviations in

Table 2: Descriptive statistics (TIMSS)

| | Pooled | Denmark | Lithuania | Norway | Spain | Sweden |
|-------------------------------------|--------|---------|-----------|--------|--------|--------|
| Student demographics | | | | | | |
| Female | 0.50 | 0.54 | 0.49 | 0.51 | 0.49 | 0.48 |
| | (0.50) | (0.50) | (0.50) | (0.50) | (0.50) | (0.50) |
| Age (months) | 122.70 | 130.55 | 128.30 | 116.59 | 117.40 | 128.76 |
| | (7.28) | (4.59) | (4.28) | (3.48) | (5.00) | (3.93) |
| Test language spoken at home | | | | | | |
| – always | 0.79 | 0.83 | 0.83 | 0.80 | 0.74 | 0.78 |
| | (0.41) | (0.37) | (0.37) | (0.40) | (0.44) | (0.41) |
| – sometimes | 0.18 | 0.16 | 0.16 | 0.18 | 0.19 | 0.21 |
| | (0.39) | (0.37) | (0.36) | (0.39) | (0.39) | (0.41) |
| – never | 0.03 | 0.01 | 0.01 | 0.01 | 0.07 | 0.01 |
| | (0.16) | (0.08) | (0.10) | (0.12) | (0.26) | (0.11) |
| Books at home | | | | | | |
| – less than one shelf (\leq =10) | 0.09 | 0.09 | 0.13 | 0.06 | 0.09 | 0.06 |
| | (0.28) | (0.28) | (0.33) | (0.24) | (0.29) | (0.24) |
| – one shelf (11-25) | 0.25 | 0.27 | 0.36 | 0.19 | 0.26 | 0.20 |
| , , , | (0.44) | (0.45) | (0.48) | (0.39) | (0.44) | (0.40) |
| – one bookcase (26-100) | 0.35 | 0.37 | 0.35 | 0.36 | 0.34 | 0.34 |
| | (0.48) | (0.48) | (0.48) | (0.48) | (0.47) | (0.47) |
| – two bookcases (101 - 200) | 0.17 | 0.16 | 0.10 | 0.20 | 0.16 | 0.22 |
| | (0.37) | (0.37) | (0.30) | (0.40) | (0.36) | (0.41) |
| – more than two bookcases (>200) | 0.14 | 0.11 | 0.07 | 0.19 | 0.15 | 0.18 |
| | (0.35) | (0.31) | (0.25) | (0.39) | (0.36) | (0.38) |
| Day covariates | | | | | | |
| Monday | 0.14 | 0.14 | 0.10 | 0.14 | 0.17 | 0.15 |
| • | (0.35) | (0.35) | (0.30) | (0.35) | (0.37) | (0.36) |
| Tuesday | 0.27 | 0.17 | 0.24 | 0.22 | 0.35 | 0.29 |
| · | (0.44) | (0.38) | (0.43) | (0.42) | (0.48) | (0.46) |
| Wednesday | 0.30 | 0.48 | 0.29 | 0.34 | 0.24 | 0.28 |
| | (0.46) | (0.50) | (0.45) | (0.47) | (0.43) | (0.45) |
| Thursday | 0.20 | 0.17 | 0.21 | 0.23 | 0.20 | 0.18 |
| • | (0.40) | (0.37) | (0.41) | (0.42) | (0.40) | (0.39) |
| Friday | 0.09 | 0.04 | 0.16 | 0.07 | 0.05 | 0.09 |
| | (0.28) | (0.19) | (0.37) | (0.25) | (0.21) | (0.28) |
| Observations | 8813 | 564 | 2116 | 2328 | 2208 | 1597 |

Notes: Own calculations for the pooled sample based on TIMSS 2011. Mean values and standard deviations (in parentheses) of the pooled and country-specific samples. The day covariates indicate the percentage of students tested on that day.

Table 3: Differences in covariates before and after the treatment (TIMSS)

| | Bef | ore | Af | ter | Befor | e - After | Normalized difference |
|-------------------------------------|--------|--------|--------|--------|---------|-----------|-----------------------|
| | Mean | (S.D.) | Mean | (S.D.) | Diff. (| P-value) | |
| Student demographics | | | | | | | |
| Female | 0.50 | (0.50) | 0.50 | (0.50) | 0.00 | (0.93) | 001 |
| Age (months) | 123.51 | (7.11) | 122.03 | (7.36) | 1.48 | (0.00) | 144 |
| Test language at home | | | | | | | |
| – always | 0.77 | (0.42) | 0.81 | (0.39) | -0.04 | (0.00) | 0.076 |
| - sometimes | 0.21 | (0.41) | 0.16 | (0.37) | 0.05 | (0.00) | 090 |
| – never | 0.02 | (0.15) | 0.03 | (0.17) | -0.01 | (0.11) | 0.024 |
| Books at home: | | | | | | | |
| – less than one shelf (\leq =10) | 0.08 | (0.27) | 0.09 | (0.29) | -0.01 | (0.03) | 0.032 |
| - one shelf (11-25) | 0.24 | (0.43) | 0.27 | (0.44) | -0.02 | (0.01) | 0.038 |
| - one bookcase (26-100) | 0.34 | (0.48) | 0.36 | (0.48) | -0.01 | (0.19) | 0.020 |
| - two bookcases (101 - 200) | 0.17 | (0.38) | 0.16 | (0.37) | 0.02 | (0.06) | 029 |
| – more than two bookcases (>200) | 0.16 | (0.37) | 0.13 | (0.33) | 0.03 | (0.00) | 069 |
| Day of test: | | | | | | | |
| - Monday | 0.13 | (0.34) | 0.15 | (0.35) | -0.01 | (0.05) | 0.029 |
| – Tuesday | 0.20 | (0.40) | 0.32 | (0.47) | -0.12 | (0.00) | 0.195 |
| – Wednesday | 0.37 | (0.48) | 0.24 | (0.43) | 0.13 | (0.00) | 203 |
| – Thursday | 0.19 | (0.39) | 0.21 | (0.41) | -0.02 | (0.00) | 0.044 |
| – Friday | 0.10 | (0.30) | 0.07 | (0.26) | 0.03 | (0.00) | 071 |
| Observations | 4000 | | 4813 | | 8813 | | |

Notes: The table reports the mean values and standard deviations of all covariates for the control group (tested before the transition into DST) and the treatment group (tested after the transition). The third column contains the difference in means and the respective p-values from testing the hypothesis that the two means are equal. The last column reports the normalized differences as suggested by Imbens and Wooldridge (2009, p. 24). Calculations based on TIMSS 2011.

Table 4: Impact of the clock change on students' performance (pooled sample)

| | (1) | (2) | (3) | (4) |
|---------------------------|-----------|--------------|------------|--------------|
| | Sample: | TIMSS (88 | 13 observ | ations) |
| a) Mathematics | | | | |
| DST effect | -4.042 | -9.131 | -3.462 | -8.139 |
| | (3.512) | (8.742) | (3.074) | (7.663) |
| b) Science | | | | |
| DST effect | -3.892 | -10.601 | -3.433 | -9.439 |
| | (3.444) | (8.586) | (2.909) | (7.293) |
| | Sample: 1 | PIRLS (132 | 255 observ | ations) |
| c) Reading | _ | | | |
| DST effect | 0.506 | 8.180 | 0.309 | 4.072 |
| | (2.695) | (6.626) | (2.306) | (5.686) |
| Sociodemographic controls | | | √ | \checkmark |
| Days & interactions | | \checkmark | | \checkmark |

Notes: Own calculations for the pooled sample based on TIMSS and PIRLS 2011. Sociodemographic controls: gender (reference: male), age (centered), age (centered, squared), books at home (reference: one bookcase), test language spoken at home (reference: always); day and interaction controls: weekday (reference: Monday), weekday×DST (reference: Monday×DST). Standard errors in parentheses. * p < 0.1, *** p < 0.05, **** p < 0.01.

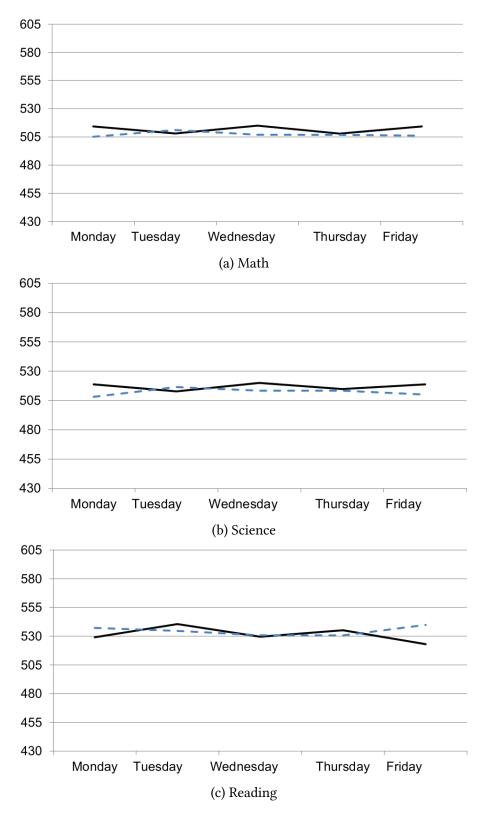


Figure 1: Performance before (solid line) and after (dashed line) the clock change over weekdays

Notes: Own calculations based on TIMSS and PIRLS 2011. The ordinate was scaled to a range of roughly one standard deviation above the average performance in reading and one standard deviation below the average performance in math.

Table 5: Students' performance by gender and by whether test language is spoken at home

| | Gen | ıder | Test lan | guage spoken | at home | | |
|----------------|---------------|---------------|---------------|------------------|--------------|--|--|
| | (1) Female | (2) Male | (3) Always | (4) Sometimes | (5) Never | | |
| | | S | ample: TI | MSS | | | |
| a) Mathematics | | | | | | | |
| DST effect | -1.969 | -3.977 | -3.556 | -4.861 | -4.450 | | |
| | (3.471) | (3.625) | (3.195) | (4.453) | (12.214) | | |
| b) Science | | | | | | | |
| DST effect | -3.391 | -2.919 | -3.359 | -5.686 | -1.980 | | |
| | (3.276) | (3.248) | (2.936) | (4.961) | (12.302) | | |
| Observations | 4393 | 4420 | 6977 | 1600 | 236 | | |
| | | Sample: PIRLS | | | | | |
| c) Reading | | | | | | | |
| DST effect | -0.303 | 1.902 | 2.268 | 2.443 | 36.075** | | |
| | (2.608) | (2.590) | (6.023) | (8.858) | (17.773) | | |
| Observations | 6595 | 6660 | 10845 | 2083 | 327 | | |

Notes: Own calculations based on TIMSS and PIRLS 2011. Regressions include the following socioeconomic controls: gender (reference: male), age (centered), age (centered), books at home (reference: one bookcase), test language spoken at home (reference: always). Standard errors in parentheses. * p < 0.1, *** p < 0.05, *** p < 0.01.

Table 6: Students' performance by books at home

| | (1) | (2) | (3) | (4) | (5) |
|----------------|---------|---------|-----------|---------|---------|
| | 0-10 | 11-25 | 26-100 | 101-200 | 200+ |
| | | Sa | mple: TIM | SS | |
| a) Mathematics | | | | | |
| DST effect | -1.804 | -0.623 | -3.076 | -4.084 | -6.001 |
| | (6.212) | (4.235) | (3.609) | (4.528) | (5.858) |
| b) Science | | | | | |
| DST effect | -3.464 | -0.231 | -2.981 | -3.900 | -5.547 |
| | (6.626) | (3.988) | (3.355) | (4.738) | (5.732) |
| Observations | 758 | 2245 | 3096 | 1456 | 1258 |
| | | Sa | mple: PIR | LS | |
| c) Reading | | | _ | | |
| DST effect | 3.066 | 4.463 | -0.255 | 0.614 | -0.388 |
| | (4.687) | (3.056) | (2.975) | (3.539) | (3.855) |
| Observations | 1048 | 3043 | 4871 | 2401 | 1892 |

Notes: Own calculations based on TIMSS and PIRLS 2011. Regressions include the following socioeconomic controls: gender (reference: male), age (centered), age (centered, squared), test language spoken at home (reference: always). Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

Table 7: Country-specific impact of the clock change on students' performance

| | (1) Denmark | (2) Finland | (3) Lithuania | (4) Norway | (5) Sweden | (6) Spain |
|----------------|----------------|----------------|------------------|---------------|---------------|--------------|
| | | | Sample: 7 | TIMSS | | |
| a) Mathematics | | | - | | | |
| DST effect | 28.546 | | 1.028 | -19.950 | -24.616 | -4.667 |
| | (25.547) | | (20.757) | (15.614) | (15.186) | (16.502) |
| b) Science | | | | | | |
| DST effect | 14.143 | | -1.307 | -14.185 | -22.154 | -6.867 |
| | (29.155) | | (21.109) | (12.003) | (14.392) | (14.249) |
| Observations | 564 | | 2116 | 2328 | 1597 | 2208 |
| | | | Sample: I | PIRLS | | |
| c) Reading | | | | | | |
| DST effect | | -10.880 | 12.194 | 15.132 | -7.796 | 11.487 |
| | | (9.435) | (16.242) | (14.067) | (24.206) | (10.140) |
| Observations | | 3502 | 2125 | 2267 | 1773 | 3588 |

Notes: Own calculations based on TIMSS and PIRLS 2011. Gender (reference: male), age (centered), age (centered, squared), books at home (reference: one bookcase), test language spoken at home (reference: always), weekday (reference: Monday), interaction weekday × DST (reference: Monday before) included as controls. Standard errors in parentheses. * p < 0.1, *** p < 0.05, **** p < 0.01.

Table 8: Impact of the clock change in the pooled sample, two weeks before and after

| | (1) | (2) | (3) | (4) |
|---------------------------|---------|--------------|----------|-------------------|
| | Samp | ole: TI | MSS (149 | 942 observations) |
| a) Mathematics | | | | |
| DST effect | -1.132 | -7.164 | -0.871 | -3.920 |
| | (2.932) | (7.230 | (2.556) | (6.286) |
| b) Science | | | | |
| DST effect | -1.757 | 10.057 | -1.552 | -6.260 |
| | (2.947) | (7.281 | (2.467) | (6.073) |
| | Sam | ole: PI | RLS (213 | 379 observations) |
| c) Reading | • | | | |
| DST effect | -1.193 | 4.324 | -1.333 | 1.524 |
| | (2.408) | (5.671 | (2.080) | (4.974) |
| Sociodemographic controls | | | ✓ | √ |
| Days & interactions | | \checkmark | | \checkmark |

Notes: Own calculations for the pooled sample based on TIMSS and PIRLS 2011. Sociodemographic controls: gender (reference: male), age (centered), age (centered), books at home (reference: one bookcase), test language spoken at home (reference: always); day and interaction controls: weekday (reference: Monday), weekday ×DST (reference: Monday before). Standard errors in parentheses. * p < 0.1, *** p < 0.05, **** p < 0.01.

Table 9: Impact of the clock change by countries, two weeks before and after

| | (1) | (2) | (3) | (4) | (5) | (6) |
|----------------|----------|---------|-----------|----------|-----------|----------|
| | Denmark | Finland | Lithuania | Norway | Sweden | Spain |
| | | | Sample: T | IMSS | | |
| a) Mathematics | | | | | | |
| DST effect | 13.378 | | 17.113 | -9.175 | -27.847** | -8.245 |
| | (21.675) | | (18.155) | (13.050) | (13.496) | (13.445) |
| b) Science | | | | | | |
| DST effect | 3.076 | | 16.361 | -9.827 | -24.831* | -8.951 |
| | (21.615) | | (17.469) | (9.808) | (13.210) | (12.154) |
| Observations | 1213 | | 3755 | 2971 | 3522 | 3481 |
| | | | Sample: I | PIRLS | | |
| c) Reading | | | | | | |
| DST effect | | -8.349 | 12.633 | 11.605 | 1.003 | 2.373 |
| | | (8.754) | (16.182) | (14.519) | (14.226) | (9.844) |
| Observations | | 4600 | 3664 | 2870 | 3458 | 6787 |

Notes: Own calculations based on TIMSS and PIRLS 2011. Gender (reference: male), age (centered), age (centered, squared), books at home (reference: one bookcase), test language spoken at home (reference: always), weekday (reference: Monday), interaction weekday × DST (reference: Monday before) included as controls. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

Table 10: Impact of the clock change on eighth-graders, one week before and after

| | (1) Pooled | (2) Finland | (3) Sweden |
|--------------|---------------|----------------|---------------|
| Math | | | |
| DST effect | -13.681 | -1.138 | -21.007* |
| | (9.203) | (14.202) | (11.514) |
| Science | , | , | , |
| DST effect | -14.707 | 2.370 | -26.008* |
| | (9.043) | (13.474) | (12.283) |
| Observations | 5591 | 3547 | 2044 |

Notes: Own calculations based on TIMSS 2011. Gender (reference: male), age (centered), age (centered), age (centered), age (centered, squared), books at home (reference: one bookcase), test language spoken at home (reference: always), weekday (reference: Monday), interaction weekday × DST (reference: Monday before) included as controls. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

A Appendix

Table A1: Plausible values of reading by country and treatment

| | Poc | Pooled | Finl | Finland | Lith | Lithuania | Norway | way | Sp | Spain | Sweden | den |
|--------------|---------|---------|---------|---------|---------|-----------|---------|---------|---------|---------|---------|---------|
| | before | after | before | after | before | after | before | after | before | after | before | after |
| - 1st PV | 538.98 | 535.80 | 565.84 | 569.35 | 531.33 | 539.15 | 511.55 | 507.23 | 524.60 | 524.21 | 540.38 | 539.63 |
| | (67.84) | (66.15) | (64.61) | (61.39) | (65.40) | (62.72) | (60.10) | (61.35) | (68.82) | (64.13) | (65.49) | (66.22) |
| - 2nd PV | 537.72 | 535.14 | 565.34 | 569.00 | 529.20 | 539.80 | -, | 507.63 | 521.80 | 521.53 | 538.90 | 542.55 |
| | (68.63) | (67.54) | (65.33) | (61.52) | (66.11) | (64.35) | (60.30) | (63.66) | | (65.18) | (66.68) | (69.01) |
| – 3rd PV | 537.57 | 535.32 | 565.11 | 568.29 | 528.65 | 538.64 | -, | 508.17 | | | 538.39 | 542.36 |
| | (68.26) | | (64.01) | (61.55) | (66.01) | (63.74) | (60.37) | (62.58) | (69.11) | (64.93) | (67.31) | (66.83) |
| – 4th PV | 538.63 | 535.07 | 565.85 | 568.09 | 529.46 | 538.76 | | 507.53 | 524.92 | 522.69 | 539.98 | 541.39 |
| | (68.18) | (67.15) | (64.82) | (61.75) | (64.49) | (64.07) | (59.94) | (63.69) | (69.53) | (64.80) | (66.78) | (68.45) |
| – 5th PV | 537.64 | 535.56 | 564.65 | 569.04 | 532.11 | 539.27 | 510.61 | 506.76 | 522.08 | 523.49 | 537.98 | 541.21 |
| | (68.32) | (66.58) | (65.18) | (62.21) | (64.37) | (63.13) | (60.30) | (61.68) | (69.36) | (64.20) | (67.05) | (67.16) |
| Observations | 6159 | 9602 | 2008 | 1494 | 702 | 1423 | 1356 | 911 | 906 | 2682 | 1187 | 586 |

Notes: Plausible values (PV) as contained in PIRLS 2011. Standard deviations in parentheses.

Table A2: Descriptive statistics (PIRLS)

| | Pooled | Finland | Lithuania | Norway | Spain | Sweden |
|-------------------------------------|--------|---------|-----------|--------|--------|--------|
| Student demographics | | | | | | |
| Female | 0.50 | 0.49 | 0.50 | 0.51 | 0.49 | 0.50 |
| | (0.50) | (0.50) | (0.50) | (0.50) | (0.50) | (0.50) |
| Age (months) | 123.70 | 129.30 | 128.35 | 116.59 | 117.43 | 128.87 |
| | (7.28) | (4.19) | (4.26) | (3.51) | (4.93) | (3.95) |
| Test language spoken at home | | | | | | |
| – always | 0.82 | 0.89 | 0.83 | 0.81 | 0.77 | 0.77 |
| | (0.39) | (0.31) | (0.37) | (0.39) | (0.42) | (0.42) |
| - sometimes | 0.16 | 0.10 | 0.16 | 0.18 | 0.17 | 0.22 |
| | (0.36) | (0.30) | (0.36) | (0.38) | (0.38) | (0.41) |
| – never | 0.02 | 0.01 | 0.01 | 0.01 | 0.06 | 0.02 |
| | (0.16) | (0.09) | (0.10) | (0.12) | (0.24) | (0.13) |
| Books at home | | | | | | |
| – less than one shelf (\leq =10) | 0.08 | 0.05 | 0.12 | 0.06 | 0.09 | 0.08 |
| | (0.27) | (0.22) | (0.33) | (0.24) | (0.29) | (0.27) |
| – one shelf (11-25) | 0.23 | 0.15 | 0.35 | 0.19 | 0.27 | 0.20 |
| | (0.42) | (0.36) | (0.48) | (0.39) | (0.44) | (0.40) |
| - one bookcase (26-100) | 0.37 | 0.42 | 0.35 | 0.37 | 0.34 | 0.34 |
| | (0.48) | (0.49) | (0.48) | (0.48) | (0.47) | (0.47) |
| - two bookcases (101 - 200) | 0.18 | 0.23 | 0.10 | 0.20 | 0.16 | 0.20 |
| | (0.39) | (0.42) | (0.30) | (0.40) | (0.37) | (0.40) |
| – more than two bookcases (>200) | 0.14 | 0.15 | 0.06 | 0.18 | 0.14 | 0.18 |
| | (0.35) | (0.36) | (0.25) | (0.39) | (0.35) | (0.38) |
| Day covariates | | | | | | |
| - Monday | 0.15 | 0.13 | 0.16 | 0.09 | 0.26 | 0.04 |
| | (0.36) | (0.33) | (0.37) | (0.29) | (0.44) | (0.19) |
| – Tuesday | 0.27 | 0.26 | 0.26 | 0.27 | 0.27 | 0.32 |
| | (0.44) | (0.44) | (0.44) | (0.45) | (0.44) | (0.46) |
| – Wednesday | 0.26 | 0.27 | 0.20 | 0.30 | 0.22 | 0.36 |
| | (0.44) | (0.45) | (0.40) | (0.46) | (0.41) | (0.48) |
| - Thursday | 0.19 | 0.23 | 0.20 | 0.23 | 0.13 | 0.17 |
| | (0.39) | (0.42) | (0.40) | (0.42) | (0.34) | (0.38) |
| – Friday | 0.13 | 0.12 | 0.17 | 0.11 | 0.13 | 0.12 |
| • | (0.33) | (0.32) | (0.38) | (0.31) | (0.33) | (0.32) |
| Observations | 13255 | 3502 | 2125 | 2267 | 3588 | 1773 |

Notes: Own calculations for the pooled sample based on PIRLS 2011. Mean values and standard deviations (in parentheses). The day covariates indicate the percentage of students tested on that day.

Table A3: Differences in covariates before and after the treatment (PIRLS)

| | Bef | ore | After | | Before - After | | Normalized difference |
|-------------------------------------|--------|--------|--------|--------|----------------|---------|-----------------------|
| | Mean | (S.D.) | Mean | (S.D.) | Diff. (P | -value) | |
| Student demographics | | | | | | | |
| Female | 0.50 | (0.50) | 0.49 | (0.50) | 0.01 | (0.45) | 009 |
| Age (months) | 124.52 | (7.15) | 123.00 | (7.32) | 1.52 | (0.00) | 148 |
| Test language at home | | | | | | | |
| - always | 0.81 | (0.39) | 0.83 | (0.38) | -0.02 | (0.02) | 0.028 |
| - sometimes | 0.17 | (0.37) | 0.15 | (0.35) | 0.02 | (0.00) | 042 |
| – never | 0.02 | (0.14) | 0.03 | (0.16) | -0.01 | (0.01) | 0.030 |
| Books at home: | | | | | | | |
| – less than one shelf (\leq =10) | 0.07 | (0.26) | 0.08 | (0.28) | -0.01 | (0.08) | 0.021 |
| - one shelf (11-25) | 0.21 | (0.41) | 0.25 | (0.43) | -0.04 | (0.00) | 0.067 |
| – one bookcase (26-100) | 0.38 | (0.48) | 0.36 | (0.48) | 0.02 | (0.05) | 024 |
| - two bookcases (101 - 200) | 0.19 | (0.39) | 0.17 | (0.38) | 0.02 | (0.00) | 036 |
| - more than two bookcases (>200) | 0.15 | (0.36) | 0.14 | (0.34) | 0.01 | (0.05) | 024 |
| Day of test: | | | | | | | |
| – Monday | 0.09 | (0.29) | 0.20 | (0.40) | -0.10 | (0.00) | 0.209 |
| – Tuesday | 0.32 | (0.47) | 0.23 | (0.42) | 0.09 | (0.00) | 139 |
| – Wednesday | 0.27 | (0.44) | 0.26 | (0.44) | 0.01 | (0.11) | 020 |
| – Thursday | 0.18 | (0.39) | 0.20 | (0.40) | -0.02 | (0.02) | 0.028 |
| – Friday | 0.14 | (0.34) | 0.12 | (0.32) | 0.02 | (0.00) | 041 |
| Observations | 6159 | | 7096 | | 13255 | | |

Notes: The table reports the mean values and standard deviations of all covariates for the control group (tested before the transition into DST) and the treatment group (tested after the transition). The third column contains the difference in means and the respective p-values from testing the hypothesis that the two means are equal. The last column reports the normalized differences as suggested by Imbens and Wooldridge (2009, p. 24). Calculations based on PIRLS 2011.

Table A4: Impact of the clock change on performance in math (pooled sample)

| | (1) | | (2) | | (3) | | (4) | |
|-----------------------------------------------------------------------------------------|-----------|---------|----------|-----------|------------|---------|-----------|----------|
| | | | | Fixed | Part | | | |
| DST effect | -4.04 | (3.51) | -9.13 | (8.74) | -3.46 | (3.07) | -8.14 | (7.66) |
| Day of test: | | | | | | | | |
| – Tuesday | | | -6.26 | (8.25) | | | -3.62 | (7.32) |
| – Wednesday | | | 0.69 | (7.97) | | | 1.87 | (7.03) |
| - Thursday | | | -6.39 | (8.20) | | | -3.73 | (7.27) |
| – Friday | | | 0.02 | (9.99) | | | 3.55 | (8.92) |
| Interaction: after \times day | | | | | | | | |
| – Tuesday \times after | | | 12.18 | (10.75) | | | 9.53 | (9.56) |
| – Wednesday \times after | | | 1.11 | (10.71) | | | 2.98 | (9.41) |
| – Thursday \times after | | | 8.10 | (11.01) | | | 7.18 | (9.73) |
| – Friday \times after | | | 0.92 | (13.99) | | | 1.76 | (12.29) |
| Student demographics | | | | | | | | |
| Female | | | | | -9.71*** | , , | | , , |
| Age (months, centered) | | | | | 0.15 | (0.19) | | (0.19) |
| Age ² (centered) | | | | | -0.06*** | (0.01) | -0.06*** | (0.01) |
| Test language at home | | | | | | | | |
| - sometimes | | | | | -10.98*** | | -10.95*** | |
| – never | | | | | -23.54*** | (5.66) | -23.46*** | (5.67) |
| Books at home: | | | | | | | | |
| -<=10 | | | | | -48.69*** | (3.77) | -48.74*** | (3.78) |
| - 11-25 | | | | | -21.19*** | (1.90) | -21.19*** | (1.90) |
| - 101-200 | | | | | 9.44*** | (2.29) | 9.46*** | (2.29) |
| ->200 | | | | | 10.88*** | (2.71) | 10.85*** | (2.70) |
| Constant | 511.92*** | (9.26) | 514.31** | * (11.14) | 529.56*** | (9.92) | 530.07*** | (11.33) |
| | | | | Randor | n Part | | | |
| σ^2 | 392.86*** | 261 88) | 386 50** | *(258 11) | 456.33***(| 300 87) | 447 90*** | (295 84) |
| σ^2 | | | | | 608.83*** | | | |
| $\begin{matrix}\sigma^2_{v_j}\\\sigma^2_{\nu_c}\\\sigma^2_{\epsilon_{ijc}}\end{matrix}$ | | | | | 3692.35*** | | | |
| Observations | 8813 | | 8813 | | 8813 | | 8813 | |
| Model F-test | 1.32 | | .49 | | 53.7 | | 33.3 | |
| p-value | .25 | | .88 | | .00 | | .00 | |

Notes: Own calculations for the pooled sample based on TIMSS 2011. Reference categories used are the following: Test language is always spoken at home, the student reports one bookcase (26-200) of books at home, and the student was tested on a Monday before the transition into DST. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

Table A5: Impact of the clock change on performance in science (pooled sample)

| | (1) | | (2) | | (3) | | (4) | |
|-----------------------------------------------------------------------------------------------|-----------|---------|----------|-----------|------------|---------|-----------|----------|
| | | | | Fixed | Part | | | |
| DST effect | -3.89 | (3.44) | -10.60 | (8.59) | -3.43 | (2.91) | -9.44 | (7.29) |
| Day of test: | | | | | | | | |
| - Tuesday | | | -5.99 | (7.90) | | | -3.00 | (6.87) |
| – Wednesday | | | 1.25 | (7.66) | | | 2.37 | (6.50) |
| – Thursday | | | -3.99 | (7.97) | | | -1.16 | (6.84) |
| – Friday | | | -0.01 | (9.51) | | | 3.06 | (8.20) |
| Interaction: after \times day | | | | | | | | |
| – Tuesday \times after | | | 14.26 | (10.72) | | | 10.83 | (9.27) |
| – Wednesday \times after | | | 4.00 | (10.55) | | | 5.75 | (8.96) |
| – Thursday \times after | | | 9.20 | (11.01) | | | 7.56 | (9.43) |
| – Friday \times after | | | 1.91 | (14.07) | | | 3.39 | (11.92) |
| Student demographics | | | | | | | | |
| Female | | | | | -7.52*** | (1.38) | | ` , |
| Age (months, centered) | | | | | 0.24 | (0.18) | | (0.18) |
| Age ² (centered) | | | | | -0.06*** | (0.01) | -0.06*** | (0.01) |
| Test language at home | | | | | | | | |
| – sometimes | | | | | -19.60*** | , , | | ` , |
| – never | | | | | -30.73*** | (4.93) | -30.69*** | (4.93) |
| Books at home: | | | | | | | | |
| -<=10 | | | | | -51.79*** | | | |
| - 11-25 | | | | | -19.87*** | | -19.87*** | |
| - 101-200 | | | | | 12.67*** | | | |
| ->200 | | | | | 14.21*** | (2.91) | 14.18*** | (2.89) |
| Constant | 517.11*** | (6.52) | 518.77** | * (8.81) | 534.58*** | (6.51) | 534.31*** | (8.15) |
| | | | | Randor | n Part | | | |
| σ^2 | 173 58*** | 121 19) | 173 72** | *(121 48) | 164.29***(| 112 89) | 164 29*** | (112 96) |
| σ^2 | | | 863.16** | | | | 533.96*** | |
| $\begin{matrix} \sigma^2_{v_j} \\ \sigma^2_{\nu_c} \\ \sigma^2_{\epsilon_{ijc}} \end{matrix}$ | | | | | 3534.70*** | | | |
| Observations | 8813 | | 8813 | | 8813 | | 8813 | |
| Model F-test | 1.28 | | .517 | | 71.5 | | 44.8 | |
| p-value | .26 | | .86 | | .00 | | .00 | |

Notes: Own calculations for the pooled sample based on TIMSS 2011. Reference categories used are the following: Test language is always spoken at home, the student reports one bookcase (26-200) of books at home, and the student was tested on a Monday before the transition into DST. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

Table A6: Impact of the clock change on performance in reading (pooled sample)

| | (1) | | (2) | | (3) | | (4) | |
|-----------------------------------------------------------------------------------------------|------------|---------|-----------|----------|-------------|---------|------------|---------|
| | | | | Fixed | Part | | | |
| DST effect | 0.51 | (2.70) | 8.18 | (6.63) | 0.31 | (2.31) | 4.07 | (5.69) |
| Day of test: | | | | | | | | |
| - Tuesday | | | 11.46* | (5.91) | | | 6.29 | (5.15) |
| – Wednesday | | | 0.48 | (6.15) | | | -1.49 | (5.31) |
| – Thursday | | | 6.12 | (6.77) | | | 2.93 | (5.77) |
| – Friday | | | -5.99 | (7.05) | | | -7.13 | (6.01) |
| Interaction: after \times day | | | | | | | | |
| – Tuesday \times after | | | -14.07* | (7.95) | | | -8.30 | (6.86) |
| – Wednesday \times after | | | -6.87 | (7.89) | | | -2.81 | (6.73) |
| – Thursday \times after | | | -12.70 | (8.81) | | | -7.36 | (7.48) |
| – Friday \times after | | | 8.60 | (9.65) | | | 8.31 | (8.18) |
| Student demographics | | | | | | | | |
| Female | | | | | 11.69*** | (1.07) | | (1.07) |
| Age (months, centered) | | | | | 0.08 | (0.13) | | (0.14) |
| Age ² (centered) | | | | | -0.07*** | (0.01) | -0.07*** | (0.01) |
| Test language at home | | | | | | | | |
| sometimes | | | | | -14.59*** | (1.76) | | (1.76) |
| – never | | | | | -34.59*** | (3.78) | -34.57*** | (3.79) |
| Books at home: | | | | | | | | |
| -<=1 | | | | | -42.55*** | (2.28) | -42.54*** | (2.28) |
| - 11-25 | | | | | -20.30*** | (1.69) | -20.25*** | (1.69) |
| - 101-200 | | | | | 12.29*** | (1.64) | 12.23*** | (1.64) |
| ->200 | | | | | 13.07*** | (2.08) | 13.00*** | (2.08) |
| Constant | 532.23*** | (9.45) | 528.09*** | (10.71) | 538.11*** | (8.38) | 536.87*** | (9.49) |
| | | |] | Randor | n Part | | | |
| $\sigma^2_{\cdot \cdot \cdot}$ | 427.16***(| 275.23) | 437.80*** | (281.92) | 330.71***(2 | 214.23) | 336.58***(| 217.96) |
| $\sigma^{2^{j}}$ | , | , | | . , | 397.54*** | , | , | , |
| $\begin{matrix} \sigma^2_{v_j} \\ \sigma^2_{\nu_c} \\ \sigma^2_{\epsilon_{ijc}} \end{matrix}$ | | | | | 3222.66*** | | | |
| Observations | 13255 | | 13255 | | 13255 | | 13255 | |
| Model F-test | 0.04 | | 1.68 | | 113 | | 69.2 | |
| p-value | 0.85 | | 0.08 | | 0.00 | | 0.000 | |

Notes: Own calculations for the pooled sample based on PIRLS 2011. Reference categories used are the following: Test language is always spoken at home, the student reports one bookcase (26-200) of books at home, and the student was tested on a Monday before the transition into DST. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

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